

**5.  $G_5$ -ESTIMATOR OF SMOOTHED NORMALIZED SPECTRAL FUNCTION OF SYMMETRIC MATRICES**

Let  $\mu_n(x)$  be a normalized spectral function of a covariance matrix  $R_m$ . The  $G_2$ -estimator for Stieltjes' transform of this function is equal to (see Section 2.2)

$$G_2(A, B, u + iv) = i \int_0^B \left\{ \frac{e^{|sp|}}{\pi} \int_{-A}^A \operatorname{Im} G_2(z) e^{-itp} dt \right\} e^{-p(v-iu)} dp, \quad v > 0.$$

Using this estimator we can try to find a consistent estimator of  $\mu_n(x)$ . But in this case two questions arise:

- 1). Will the estimator  $G_2$  be equal to Stieltjes' transform of a distribution function?
- 2). The spectral function  $\mu_n(x)$  may have a discontinuity. Therefore it is very difficult to use the inverse formula for Stieltjes' transform for finding  $\mu_n(x)$ , using the  $G_2$ -estimator. To overcome these difficulties we can use the so-called smoothed normalized spectral functions

$$\mu_n(x, \varepsilon) = \frac{1}{\pi} \int_{-\infty}^x \operatorname{Im} \frac{1}{m_n} \operatorname{Tr} [R_m - (y + i\varepsilon)]^{-1} dy, \quad \varepsilon > 0.$$

It can be shown that

$$\mu_n(x, \varepsilon) = \frac{1}{\pi} \int_{-\infty}^{\infty} \frac{\mu_n(x + \varepsilon y)}{1 + y^2} dy, \quad \varepsilon > 0.$$

Therefore we call  $\mu_n(x, \varepsilon)$  a smoothed normalized spectral function. Consider the estimator

$$G_5(A, B, x, \varepsilon) = \frac{1}{\pi} \int_{-\infty}^x \operatorname{Im} G_2(A, B, y + i\varepsilon) dy, \quad \varepsilon > 0.$$

It is easy to prove that under the conditions of Theorem 2.1 such estimator  $G_5$  of  $\mu_n(x, \varepsilon)$  is consistent: with probability one, for any  $\varepsilon > 0$  and  $x$

$$\lim_{B \rightarrow \infty} \lim_{A \rightarrow \infty} \lim_{n \rightarrow \infty} \{G_5(A, B, x, \varepsilon) - \mu_n(x, \varepsilon)\} = 0.$$