## 14. $G_{14}$ -estimator of regularized discriminant function

If matrix R is singular or ill-conditioned, then instead of the Mahalanobis distance  $\alpha$ , its regularized analog is considered

$$\alpha_{\varepsilon} = (\vec{a}_1 - \vec{a}_2)^T [\varepsilon I + R]^{-1} (\vec{a}_1 - \vec{a}_2), \ \varepsilon > 0.$$

The regularized distance has more useful properties than the distance  $\alpha$ . To prove asymptotic normality of *G*-estimators, it is not necessary that the random vectors  $\vec{\xi}$  and  $\vec{\eta}$  be normally distributed. As was mentioned in the previous chapters, the estimator

$$G_{\varepsilon} = \left(\hat{\vec{a}}_1 - \hat{\vec{a}}_2\right)^T \left[\varepsilon I + \hat{R}\right]^{-1} \left(\hat{\vec{a}}_1 - \hat{\vec{a}}_2\right), \ \varepsilon > 0,$$

with empirical mean vectors and the covariance matrix  $\hat{\vec{a}}_1$ ,  $\hat{\vec{a}}_2$ ,  $\hat{R}$ , is inappropriate for solving multivariate classification problems. Indeed, with the increase of m, the number of components of the vectors  $\vec{\xi}$  and  $\vec{\eta}$ , the number of observations needed for obtaining a given accuracy in the Mahalanobis distance estimation grows rapidly. In this section we assert that under some conditions, there exists an asymptotically Normal *G*-estimator for the regularized discriminant function, provided that

$$\lim_{n_1,n_2 \to \infty} \sup \left[ mn_1^{-1} + mn_2^{-1} \right] < \infty.$$

Let  $\vec{x}_i, \ \vec{y}_j; \ i = 1, \dots, n_1; \ j = 1, \dots, n_2$  be independent observations of *m*-dimensional independent random vectors  $\vec{\xi}$  and  $\vec{\eta}$  respectively. We call the expression

$$G_{14}\left(\vec{x}\right) = \left\{\vec{x} - \frac{1}{2}\left(\hat{\vec{a}}_{1} - \hat{\vec{a}}_{2}\right)\right\}^{T} \left[\varepsilon I + \varepsilon \theta_{n_{1},n_{2}}^{-1}\hat{R}\right]^{-1} \left(\hat{\vec{a}}_{1} - \hat{\vec{a}}_{2}\right) \\ - \left[n_{1}^{-1} + n_{2}^{-1}\right]\varepsilon \theta_{n_{1},n_{2}}^{-1} \operatorname{Tr} \hat{R} \left[\varepsilon I + \varepsilon \theta_{n_{1},n_{2}}^{-1}\hat{R}\right]^{-1}$$

the  $G_{14}$ -estimator for the regularized discriminant function. Here  $\theta_{n_1,n_2}$  is the nonnegative solution of the equation [Gir54, p.615]

$$1 - k_{n_1, n_2} + k_{n_1, n_2} \theta_{n_1, n_2}^{-1} \operatorname{Tr} \hat{R} \left[ \theta_{n_1, n_2} I + \hat{R} \right]^{-1} = \varepsilon \theta_{n_1, n_2},$$
  
$$k_{n_1, n_2} = m \left[ n_1 + n_2 - 2 \right]^{-1}$$

It can be seen that there exists a unique nonnegative solution of this equation.

THEOREM 14.1. [Gir54, p.615] Let the conditions of Theorem 13.1 be satisfied. Then

$$\lim_{n_1, n_2 \to \infty} \max_{i=1,2} \mathbf{P} \left\{ \left[ G_{14} \left( \vec{\xi_i} \right) - \left\{ \vec{\xi_i} - \frac{1}{2} \left( \vec{a_1} - \vec{a_2} \right) \right\}^T \left( \varepsilon I + R \right)^{-1} \left( \vec{a_1} - \vec{a_2} \right) \right] \\ \times \sqrt{\frac{n_1 + n_2 - 2}{V_m}} < x \right\} = \frac{1}{\sqrt{2\pi}} \int_{-\infty}^x \exp\left\{ -y^2/2 \right\} \mathrm{d}y,$$

where  $\vec{\xi}_i$  is an observation which does not depend on  $\vec{x}_i$ ,  $\vec{y}_j$ ;  $i = 1, ..., n_1$ ;  $j = 1, ..., n_2$ , distributed as  $N\{\vec{a}_1, R\}$  or  $N\{\vec{a}_2, R\}$  and  $V_m$  are certain constants.

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